

# Book of Abstracts

# 15<sup>th</sup> European Young Statisticians Meeting



September 10–14, 2007  
Castro Urdiales, Spain

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## Welcome to 15th European Young Statisticians Meeting Castro Urdiales (Spain)

On behalf of the Local Organizing Committee I am pleased to welcome all of you to Castro Urdiales. The European Young Statisticians Meetings are held every two years. The idea of the meeting is to provide young researchers, perhaps just started in, or about to enter post-doctoral positions (and reasonably committed to remaining in research), an introduction to the international scene within the broad subject area. The venue of the European Young Statisticians Meeting is for the first time in Spain.

The contributions of the meeting keep a healthy balance between the theoretical and practical aspects of probability and statistics. All the talks are organized accordingly to 12 different topics in 17 sessions. There are no parallel sessions. In addition, there are two keynote lectures delivered by important scientists to which I want to thank their acceptance to participate in our Meeting.

On the other hand, the programme also offers cultural events that will allow us to discover the cities, gastronomy and sights of the dynamic and historical Spanish Communities of Cantabria and Basque Country as well as the charm and hospitality of their people.

I also would like to thank to all people that have actively participated in organizing the conference and to those entities and public administrations that have provided financial support. Particular I want to express my acknowledgement to Professor Laureano González Vega, Director of the International Center of Mathematical meetings, for his inestimable help and worthy advice.

I wish you an enjoyable and successful meeting in Castro Urdiales, hoping that you will find the conference and the city a suitable setting to share your ideas and time.

Inés M<sup>a</sup> del Puerto  
Chair of the 15th EYSM



## Schedule

	Monday, 10	Tuesday, 11	Wednesday, 12	Thursday, 13	Friday, 14
9:05-9:30	Registration	Presentation	Presentation	Presentation	Presentation
9:30-9:55	Opening	Presentation	Presentation	Presentation	Presentation
9:55-10:20	Keynote	Presentation	Presentation	Keynote	Presentation
10:20-10:45	Lecture	Presentation	Presentation	Lecture	Presentation
10:45-11:05	Coffee Break	Coffee Break	Coffee Break	Coffee Break	Closing
11:05-11:30	Presentation	Presentation	Presentation	Presentation	
11:30-11:55	Presentation	Presentation	Presentation	Presentation	
11:55-12:20	Presentation	Presentation	Presentation	Presentation	
12:20-12:45	Presentation	Presentation	Presentation	Presentation	
12:45-15:25	Lunch	Lunch	Lunch <sup>1</sup>	Lunch	Lunch <sup>2</sup>
15:25-15:50	Presentation	Presentation	Excursion to Guggenheim museum in Bilbao	Presentation	
15:50-16:15	Presentation			Presentation	
16:15-16:35	Coffee Break	Coffee Break		Coffee Break	
16:35-17:00	Presentation	Presentation		Presentation	
17:00-17:25	Presentation	Presentation		Presentation	
17:25-17:50	Presentation	Presentation		Presentation	
18:00-20:00	Guided Visit to Castro Ur- diales				
20:00-	Welcome Drink			Closing Dinner	

<sup>1</sup> This day the lunch will be until 14:50, due to the excursion will begin at 15:00.

<sup>2</sup> It can be chosen to have lunch on Sunday, 9 or Friday, 14.

Monday, September 10	
	Keynote Lecture
9:55-10:45	<b>Daniel Peña</b> New Approaches for Analyzing Heterogeneity in Statistical Models
	Session 1: Hypothesis Testing I Chairman: Juan Romo
11:05-11:30	<b>Pierpaolo Brutti</b> On an Adaptive Goodness-of-Fit test with Finite Sample Validity for Random Design Regression Models
11:30-11:55	<b>Fragiadakis Kostas</b> Test of fit for symmetric variance gamma distributions
11:55-12:20	<b>Stefanie Hayoz</b> Behavior of nonparametric tests in longitudinal designs
12:20-12:45	<b>Gery Geenens</b> Nonparametric test for conditional independence in two-way contingency tables
	Session 2: Stochastic Models in Finance I Chairman: Petra Posedel
15:25-15:50	<b>Carl Lindberg</b> Robust portfolio optimization
15:50-16:15	<b>Ignacio Cascos</b> The financial risk of a set portfolio
	Session 3: Stochastic Processes I Chairman: Alfonso Ramos
16:35-17:00	<b>Amir Aliev</b> On the principle of smooth fit in optimal stopping problems
17:00-17:25	<b>Erika Fülöp</b> Statistical problems in a discrete time random field HJM type interest rate model
17:25-17:50	<b>Lars Andersen</b> Loss Rate Asymptotics

Tuesday, September 11	
	Session 4: Bayesian Statistics Chairman: Pierpaolo Brutti
9:05-9:30	<b>Roxana Ciumara</b> Convergence rate results for some empirical Bayes estimators
9:30-9:55	<b>Evgeny Burnaev</b> Quickest detection of intensity change for Poisson process in generalized Bayesian setting
9:55-10:20	<b>Matteo Ruggiero</b> Bayesian Nonparametric Construction of Fleming-Viot Models in Population Genetics
10:20-10:45	<b>Hannes Müller</b> Bayesian Transgaussian Kriging
	Session 5: Stochastic Models in Finance II Chairman: Ignacio Cascos
11:05-11:30	<b>Petra Posedel</b> Asymptotic analysis for a simple explicit estimator in Barndor-Nielsen and Shephard stochastic volatility models
11:30-11:55	<b>Ildikó Vitéz</b> Location as risk factor Spatial analysis of an insurance data-set
11:55-12:20	<b>Pavlo Shelyazhenko</b> Bounded arbitrage for multi-period model of financial market with discrete time
12:20-12:45	<b>Ivan Mitov</b> A discrete model for correlated default times and its application to CDO
	Session 6: Stochastic Processes II Chairman: Manuel Mota
15:25-15:50	<b>Iryna Rozora</b> Gaussian Process Simulation with application of the Theory of Square-Gaussian Processes
15:50-16:15	<b>Jaroslav Sevcik</b> Repairable systems with general repair
	Session 7: Resampling Chairman: Paulo Canas Rodrigues
16:35-17:00	<b>Claudia Kirch</b> Resampling methods for the change analysis of dependent data

Tuesday, September 11	
	Session 8: Density Estimation Chairman: Miguel González
17:00-17:25	<b>José E. Chacón</b> Bootstrap Bandwidth Selection Using an h-Dependent Pilot Bandwidth
17:25-17:50	<b>Julia Dony</b> Uniform consistency of kernel type estimators and conditional U statistics with general bandwidths

Wednesday, September 12	
	Session 9: Statistical Applications I Chairman: Andrew Parnell
9:05-9:30	<b>Patricia Geli</b> Modeling the Mechanism of Postantibiotic Effect
9:30-9:55	<b>Emma Holian</b> Mixture-Regression Cluster Model applied to Longitudinal Microarray Experiments
9:55-10:20	<b>Corinne Dahinden</b> Graphical Modeling for Discrete Random Variables with Application to Tissue Microarray (TMA) Experiments
10:20-10:45	<b>Robert Breitnecker</b> Analysing Regional Firm Startup Activity Using Geographically Weighted Regression: The case of Austria
	Session 10: Markov Processes I Chairman: Omiros Papaspiliopoulos
11:05-11:30	<b>Alina Nicolae</b> On asymptotic behavior of a finite Markov chain
11:30-11:55	<b>Sebastian Müller</b> Criteria for transience of branching Markov chains
11:55-12:20	<b>Krzysztof Latuszynski</b> A Regeneration Proof of the Central Limit Theorem for Uniformly Ergodic Markov Chain

Wednesday, September 12	
	Session 10: Markov Processes I Chairman: Omiros Papaspiliopoulos
12:20-12:45	<b>Vessela Stoimenova</b> On the statistics of branching processes with a random number of ancestors

Thursday, September 13	
	Session 11: Extreme Value Theory Chairman: Chen Zhou
9:05-9:30	<b>Marta Ferreira</b> Asymptotic and pre-asymptotic tail behaviour of a power max-autoregressive model
9:30-9:55	<b>Jan Dienstbier</b> Estimators of the extreme value index based on quantile regression
	Keynote Lecture
9:55-10:45	<b>Juan Antonio Cuesta Albertos</b> Random projections and goodness of fit tests for multidimensional data
	Session 12: Markov Processes II Chairman: Sebastian Müller
11:05-11:30	<b>Andrew Golightly</b> MCMC Sampling for Diffusion Processes
11:30-11:55	<b>Omiros Papaspiliopoulos</b> Recent advances in the simulation and inference for stochastic differential equation models
11:55-12:20	<b>Julie Lyng Forman</b> Estimating equations and inference from diffusion driven models
12:20-12:45	<b>Evangelos E. Vassiliou</b> Examples on lag distributed models subject to nonnegative divided differences of orders 2, 3 and 4

Thursday, September 13	
	Session 13: Hypothesis Testing II Chairman: José E. Chacón
15:25-15:50	<b>Jan Somorcik</b> Some multivariate multi-sample tests based on spatial medians
15:50-16:15	<b>Piotr Majercki</b> Approximations to Most Powerful Invariant Tests for Multinormality
	Session 14: Random Matrices Chairman: Inés del Puerto
16:35-17:00	<b>Guangming Pan</b> Large dimensional random matrices
	Session 15: Time Series Chairman: Inés del Puerto
17:00-17:25	<b>Paulo Canas Rodrigues</b> Principal Component Analysis of Dependent Data
17:25-17:50	<b>Martina Hancova</b> Comparison of prediction quality of the best linear unbiased predictors in time series linear regression models

Friday, September 14	
	Session 16: Statistical Applications II Chairman: Rodrigo Martínez
9:05-9:30	<b>Andrew Parnell</b> Recent Advances in Palaeoclimate Reconstruction
9:30-9:55	<b>Chen Zhou</b> On spatial extremes: with application to a rainfall problem
	Session 17: Theory of Codes Chairman: Maroussia Bojkova
9:55-10:20	<b>Aleksandra Popovska</b> Random Codes
10:20-10:45	<b>Natasha Ilievska</b> Probabilistic models in quasigroup error-detecting codes

# Keynote Lectures



## New approaches for analyzing heterogeneity in statistical models

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### Abstract

New procedures for detecting heterogeneity in a sample with respect to a given class of models are presented. The relationship among the problems of outlier detection, cluster analysis, discriminant analysis and model selection is reviewed. It is shown that similar ideas can be used to solve this problems. Different approaches will be presented and compared in examples and Monte Carlo studies

## Random projections and goodness of fit tests for multidimensional data

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Universidad de Cantabria, Spain.

### Abstract

Counterexamples showing two different multidimensional distributions with a common marginal are very well known. The usual way to construct those counterexamples starts by fixing the marginal and, then, constructing two different distributions sharing this marginal. In [1] a different point of view is taken. There, the authors begin by having two multidimensional distributions  $P$  and  $Q$ , and, then, they consider the following problem: Given a continuous probability measure  $\mu$  (for instance, gaussian), which is the  $\mu$ -measure of the vectors,  $h$ , which satisfy that the (one-dimensional) marginals of  $P$  and  $Q$  along the line determined by  $h$  coincide? The answer is 1 if  $P = Q$  and 0 if  $P$  and  $Q$  are different. Two sample goodness-of-fit tests follow straightforward from this result. Moreover, it was shown in [2] that this result can be extended to cover some families of distributions, thus providing ways to construct goodness-of-fit tests to those families. In particular, there it is shown that a distribution is gaussian if and only if almost every (one-dimensional) projection is gaussian. In this talk I will comment those results and will present some applications to real data sets.

### References

- [1] Cuesta-Albertos, J.A., R. Fraiman and T. Ransford (2007). *A sharp form of the Cramér-Wold theorem*. To appear in J. Theoret. Probab.
- [2] Cuesta-Albertos, J.A., E. del Barrio, R. Fraiman and C. Matrán (2007). *The random projection method in goodness of fit for functional data*. To appear in Comput. Statist. and Data Anal.

Session 1

# Hypothesis Testing I



## On an Adaptive Goodness-of-Fit test with Finite Sample Validity for Random Design Regression Models

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**Keywords:** Goodness-of-fit. Nonparametric Regression. Adaptive test. Random Design. Separation Rates. Warped Wavelets. U-statistics. Spherical needlets.

**AMS:** 62G08, 62G10

### Abstract

Given an i.i.d. sample  $\{(X_i, Y_i)\}_{i \in \{1, \dots, n\}}$  from the random design regression model  $Y = f(X) + \epsilon$  with  $(X, Y) \in [0, 1] \times [-M, M]$ , we consider the problem of testing the (simple) null hypothesis “ $f = f_0$ ”, against the alternative “ $f \neq f_0$ ” from a non-asymptotic point of view and for a fixed  $f_0 \in L^2([0, 1], G_X)$ , where  $G_X(\cdot)$  denotes the (known) marginal distribution of the design variable  $X$ . The procedure proposed is an adaptation to the regression setting of a multiple testing technique introduced by Fromont and Laurent [2], and it amounts to consider a suitable collection of unbiased estimators of the  $L^2$ -distance  $d_2(f, f_0) = \int [f(x) - f_0(x)]^2 dG_X(x)$ , rejecting the null hypothesis when at least one of them is greater than its  $(1 - u_\alpha)$  quantile, with  $u_\alpha$  calibrated to obtain a level- $\alpha$  test. These estimators are built upon the *warped wavelet* system recently introduced by Picard and Kerkyacharian [1], and the resulting testing procedure turns out to be adaptive over a particular collection of approximation spaces linked to the classical Besov spaces. Possible extensions of the proposed procedure to other settings (e.g. two-sample problems, unknown design density, composite-hypotheses), and an analogous goodness-of-fit test on the sphere based on a new type of spherical wavelets, called needlets (see [3]) will also be sketched.

### References

- [1] Kerkyacharian, G. and Picard D. (2004). *Regression in random design and warped wavelets*. Bernoulli, 10(6), 1053–1105.
- [2] Fromont M. and Laurent B. (2006) *Adaptive goodness-of-fit tests in a density model*. The Annals of Statistics, 34(2), 680–720.
- [3] Narcowich, F., Petrushev P. and Ward J. (2006). *Decomposition of Besov and Triebel-Lizorkin spaces on the sphere*. J. Funct. Anal., 238, 530–564.

## Test of fit for symmetric variance gamma distributions

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**Keywords:** Variance Gamma Distribution, Characteristic Function, EM - type algorithm, Goodness of Fit

### Abstract

The aim of this paper is to provide goodness-of-fit tests for the symmetric normal variance gamma distribution (SNVG). The NVG distribution is defined as a mixture of a normal distribution with a gamma distribution. The SNVG distribution is a three - parameter model, denoted by  $SNVG(\delta, c, \lambda)$  which depends on a location parameter  $\delta \in \mathbf{R}$ , a scale parameter  $c > 0$ , and a shape parameter  $\lambda > 0$ .

Suppose that on the basis of independent copies  $X_1, X_2, \dots, X_n$ , of a random variable  $X$  we wish to test the null hypothesis

$H_0$ : The law of  $X$  is  $SNVG(\delta, c, \lambda)$  for some  $\delta \in \mathbf{R}$ ,  $\sigma > 0$  and  $\lambda > 0$ .

We study a new family of omnibus tests of  $H_0$  based on the empirical characteristic function (CF). Despite the fact that the density function of  $X$  is complicated, the CF,  $\phi(t) = \mathbf{E}(e^{itX})$  of  $X$  is simply

$$\phi(t; \delta, c, \lambda) = e^{i\delta t}(1 + c^2 t^2)^{-\lambda}. \quad (1)$$

Its main characteristics are that the empirical CF is an unbiased and consistent estimator of the CF population and that, under specific sampling situations - including the present one - is more convenient to employ methods based on it, rather than classical methods such as moment based or likelihood inference methods. The "Fourier approach" has been proved to be a simple and very powerful tool for statistical inference, particularly in goodness - of - fit problems. Many inference procedures based on the empirical characteristic function have appeared lately in the literature.

A critical issue is the estimation of the parameters. Due to the complexity of the density function a direct maximization of the likelihood function is difficult while the method of moments estimates do not converge rapidly. In order to exceed this problem we use an EM - type algorithm.

The results of a Monte Carlo study would be presented. The Monte Carlo study was implemented by drawing 1000 samples of size  $n = 50$ ,  $n = 100$  and  $n = 200$ . However, since the null distribution of the test statistic depends on the value of the parameter  $\lambda$ , which is unknown, we resort to a parametric bootstrap procedure.

## Behavior of nonparametric tests in longitudinal designs

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Institute of Mathematical Statistics and Actuarial Science.  
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**Keywords:** Nonparametric procedures. Longitudinal data. Repeated measurements. Simulation study.

**AMS:** 62G10, 62G35

### Abstract

Nonparametric procedures need less model assumptions than the analogous parametric ones. Moreover the nonparametric procedures are more robust against deviations from the model assumptions. It is well known that the Friedman test is used to analyze the treatment effect in the randomized complete block design (see e.g. [3]). The Friedman test could also be used in a longitudinal design to analyze the time effect. Alternatives are an ANOVA type test (see e.g. [1]) and the parametric ANOVA test for repeated measures (see e.g. [2]). The three tests are compared with respect to the power and the control of the nominal level by means of a simulation study. This study incorporates different covariance structures are used to model the dependence among the repeated measurements and the results are based on 40'000 replications for each situation considering small sample sizes.

### References

- [1] Brunner, E. , Domhof, S. and Langer, F. (2002). *Nonparametric analysis of longitudinal data in factorial experiments*. Wiley, New York.
- [2] Davis, C. S. (1952). *Statistical methods for the analysis of repeated measurements*. Springer, New York.
- [3] Lehmann, E. L. (1975). *Nonparametrics: Statistical methods based on ranks*. Holden-Day, Inc., San Francisco.

## Nonparametric test for conditional independence in two-way contingency tables

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**Keywords:** contingency tables; conditional independence; chi-square test; nonparametric regression; long-range dependent process.

**AMS:** 62H17, 62E20, 62G08, 62G20, 62M07.

### Abstract

Consider a two-way contingency table, built on two categorical variables  $R$  and  $S$ . A fundamental question in this context is whether  $R$  and  $S$  are independent or not. This over-studied testing problem mostly relies on the chi-square or the likelihood ratio test statistics. The principal drawback of this classical way of doing is that probabilities for an individual to fall in a cell of the table are assumed to be equal from one individual to another, so that we are not treating each individual as such but rather a group of supposed homogeneous units. This is often highly unrealistic, since in most of the practical situations some possibly known characteristics of each individual ought to influence or be associated with  $R$ ,  $S$ , or both, and therefore influence the whole dependence structure of the table. In this case, a more judicious idea seems to analyze the conditional joint distribution of  $R$  and  $S$  given the vector of covariates, say  $X$ , and then to test for the conditional independence between  $R$  and  $S$  given  $X$ . In the literature, a striking fact is that the estimation of conditional probabilities associated with categorical responses, given a vector of covariates, is almost always treated via logistic regression methods, most of the time with very few validation of this parametric assumption. In this talk, we first present a nonparametric estimation procedure for the conditional probabilities to fall in each cell of the table. These estimates can be used as such, or be employed to validate a parametric assumption, like the logistic one. Secondly, we propose a generalization of the chi-square and the likelihood ratio tests to the case of conditional independence test, based on the above-mentioned nonparametric estimates of the conditional probabilities. The asymptotic law of the proposed test statistics is derived, and a simulation study illustrates the performance of the procedure in finite-sample situations.

**Acknowledgements:** Financial support from the "Interuniversity Attraction Pole", Phase VI from the Belgian Science Policy is gratefully acknowledged.

## Session 2

# Stochastic Models in Finance I



## Robust portfolio optimization

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**Keywords:** Black-Scholes model, robust portfolio optimization, Markowitz' problem,  $1/n$  strategy, ranks.

**AMS:** 60H30,62P05

### Abstract

It is widely recognized that when classical optimal strategies are used with parameters estimated from data, the resulting portfolio weights are remarkably volatile and unstable over time. The predominant explanation for this is the difficulty to estimate expected returns accurately. We propose to parameterize an  $n$  stock Black-Scholes model as an  $n$  factor Arbitrage Pricing Theory model where each factor has the same expected return. Hence the non-unique volatility matrix determines both the covariance matrix and the expected returns. This enables the investor to impose views on the future performance of the assets in the model. We derive an explicit strategy  $\pi^*$  which solves Markowitz' continuous time portfolio problem in our framework. The optimal strategy is to implicitly keep  $1/n$  of the wealth invested in stocks in each of the  $n$  underlying factors. To illustrate the long-term performance of  $\pi^*$ , we apply it out-of-sample to a large data set. We find that it is stable over time and outperforms all the underlying market assets in terms of Sharpe ratios. Further,  $\pi^*$  had a significantly higher Sharpe ratio than the classical  $1/n$  strategy.

## The financial risk of a set portfolio

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**Keywords:** Coherent risk measure. Random set. Risk measure. Stochastic ordering.  
**AMS:** 91B30, 91B82, 60D05

### Abstract

We model a risky portfolio as a random set in the  $d$ -dimensional Euclidean space. A slight modification of the definition of risk measures for vector portfolios given by Cascos and Molchanov [4] makes it valid for the new set portfolios. In this framework, the risk of each set portfolio is given in terms of a subset of  $\mathbb{R}^d$ . Some relations between risk measures for set portfolios and stochastic orderings for random sets will be briefly discussed.

**Acknowledgements:** The financial support of the Spanish Ministry of Education and Science under grant MTM2005-02254 is acknowledged.

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Session 3

# Stochastic Processes I



## On the principle of smooth fit in optimal stopping problems

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**Keywords:** optimal stopping, smooth fit, free-boundary problems

**AMS:** 60G40

### Abstract

Given a function  $G: \mathbb{R}^n \rightarrow \mathbb{R}$  and a Markov process  $(X_t)_{t \geq 0}$  consider the optimal stopping problem  $V(x) = \sup_{\tau} \mathbf{E}_x G(X_{\tau})$ . It is well known that the optimal stopping time is given by  $\tau^* = \inf\{t : X_t \in D\}$  where the set  $D$  is given by  $D := \{x : V(x) = G(x)\}$  and  $V$  is the smallest superharmonic function that dominates  $G$ . This characterization could be reformulated as the free-boundary problem, where both  $V$  and  $D$  have to be determined. To get the explicit solution of the arising differential equations one has to impose additional boundary conditions on  $V$ . Often, these conditions follow from the principle of smooth fit that states  $V' = G'$  on  $\partial D$  (in one-dimensional case) under various assumptions on  $G$  and  $X$ . These results are well-known and widely used, but only the weakest of them could be extended directly to the multi-dimensional case due to topological complications. Therefore, "smooth fit" is considered in a topology weaker than the standard topology of  $\mathbb{R}^n$ . In the article sufficient conditions are presented for this principle of smooth fit to hold in the multi-dimensional case.

**Acknowledgements:** The author is grateful to professor Shiryaev for the formulation of the problem.

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## Statistical problems in a discrete time random field HJM type interest rate model

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**Keywords:** HJM type forward interest rate model, strong consistency, LAN.

### Abstract

We consider some statistical questions arising in a discrete time Heath-Jarrow-Morton (HJM) type forward interest rate model, where the interest rate curves are driven by a geometric spatial autoregression field. Such models were proposed by Gáll, Pap and Zuijlen [2].

Our aim is to test the autoregression parameter. We study strong consistency of the maximum likelihood estimator of the parameter  $\rho$ . The difficulty is that the underlying sample consists of nonindependent random variables. Moreover, no explicit formula is available for the maximum likelihood estimators of  $\rho$ .

In the stable ( $|\rho| < 1$ ) and unstable ( $|\rho| = 1$ ) cases we showed local asymptotic normality (LAN) of the sequence of the related statistical experiments in [1] in the sense of Le Cam [3], see also Van der Vaart [4]. The main gain of this result is that we obtain at once asymptotically optimal tests.

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## Loss Rate Asymptotics

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**Keywords:** Lévy process, reflection, loss rate, asymptotics, heavy tail, asymptotics.  
**AMS:** 60G51 62P30

### Abstract

We consider a Lévy process  $\{S_t\}$  which is reflected at 0 and  $K > 0$ . The reflected process  $\{V_t^K\}$  is given as a solution to a Skorokhod problem, which implies a representation  $V_t^K = V_0 + S_t + L_t^0 - L_t^K$ , where  $\{L_t^0\}$  and  $\{L_t^K\}$  are the local times at 0 at  $K$ , respectively. The regenerative structure of  $\{V_t^K\}$  yields a stationary distribution denoted  $\pi_K$  and the loss rate is defined as the mean of  $L_1^K$  in the stationary situation. The loss rate was studied in Asmussen & Pihlsgård [1], where it was expressed in terms of the characteristic triplet of  $\{S_t\}$  and  $\pi_K$ , and asymptotics of the loss rate as  $K \rightarrow \infty$  was derived in the case of negative drift and light tails. Asymptotics for positive drift is straightforward by reversing the role of the barriers 0 and  $K$  and using a conservation law.

We use the expression for the loss rate from [1] to derive asymptotics in the case of negative drift and heavy tails, as well as in the case of zero drift. In the zero drift case, functional limit theorems (with a Brownian or stable limit) play an important role and are based on continuity properties of the loss rate.

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Session 4

# Bayesian Statistics



## Convergence rate results for some empirical Bayes estimators

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**Keywords:** Empirical Bayes. Asymptotic optimality. Convergence rate.

**AMS:** 62C12

### Abstract

The problem of asymptotic optimality and convergence rates for empirical Bayes estimators of some parameters was studied by different authors. Using a quadratic loss function, Liang [2] proved that, under certain conditions, the empirical Bayes estimator of the scale parameter in Pareto distribution is asymptotically optimal. Huang and Liang [1] analyzed the empirical Bayes estimator of the truncation parameter of truncated distribution family under asymmetric Linex loss. Later, Shi et al. [3] derived the convergence rate of empirical Bayes estimator for two-dimensional truncation parameters under Linex loss.

In this paper we evaluate the empirical Bayes estimators of the parameters of some truncated-type distributions, considering different classes of loss functions. We find the Bayes estimator and empirical Bayes estimator for the parameter taken into account. We establish conditions for asymptotic optimality and derive the associated rate of convergence to optimal risk. Finally, we consider a special case when the prior distribution is of a certain form.

**Acknowledgements:** The research is supported through CNCSIS Project 83/2007

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## Quickest detection of intensity change for Poisson process in generalized Bayesian setting

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**Keywords:** Poisson disorder problem. Quickest detection. Optimal stopping.

**AMS:** 60G40, 62M20

### Abstract

We observe a stochastic process  $X_t = \int_0^t I(s < \theta) dN_s^{\lambda_0} + \int_0^t I(s \geq \theta) dN_{s-\theta}^{\lambda_1}$ ,  $X_0 = 0$ , where  $N_t^{\lambda_0}$  and  $N_t^{\lambda_1}$  are Poisson processes with known intensities  $\lambda_0$  and  $\lambda_1$ ,  $\lambda_0 > \lambda_1$ ,  $I(\cdot)$  is an indicator function. Nonrandom unknown parameter  $\theta \in [0, \infty]$  is a time when the "disorder" appears. Let  $\tau = \tau(\omega)$  be a finite stopping time with respect to the filtration generated by the process  $X_t$ . We interpret  $\tau$  as the decision that the disorder has happened at the time  $\tau(\omega)$ . Let  $P_t$  be the distribution of the process  $X_t$  under the assumption that the disorder happened at the time  $\theta = t$ . For every  $T > 0$  we denote by  $\mathcal{M}_T$  the set of stopping times with the mean time  $T = E_\infty \tau$  until the false alarm. For  $\tau \in \mathcal{M}_T$  introduce the risk  $B(T; \tau) = \frac{1}{T} \int_0^\infty E_\theta (\tau - \theta)^+ d\theta$ . The stopping time  $\tau_T^* \in \mathcal{M}_T$  is called optimal if  $B(T) = B(T; \tau_T^*) = \inf_{\tau \in \mathcal{M}_T} B(T; \tau)$ . Let  $\psi_t$  be the Shiryaev's process,  $\psi_t = \int_0^t \frac{L_t}{L_s} ds$ , where  $L_t = \exp\left(\log\left(\frac{\lambda_1}{\lambda_0}\right) X_t - (\lambda_1 - \lambda_0)t\right)$  is the likelihood process. Denote  $\rho = \frac{1}{2} \left(\frac{\lambda_1 - \lambda_0}{\sqrt{\lambda_0}}\right)^2$ ,  $\beta = \frac{1}{6} \left(\frac{\lambda_1 - \lambda_0}{\sqrt{\lambda_0}}\right)^3$ ,  $(-Ei(-x)) = \int_x^{+\infty} \frac{e^{-t}}{t} dt$  ( $x > 0$ ) – integral exponential function.

**THEOREM.** For any  $T > 0$  an optimal stopping time  $\tau_T^*$  exists in  $\mathcal{M}_T$  such that  $\tau_T^* = \inf\{t \geq 0 : \psi_t \geq T\}$ . Moreover, the generalized Bayesian risk  $B(T) = \frac{V(0; T)}{T}$ , where  $V(x; T)$  solves differential-difference equation  $(1 - (\lambda_1 - \lambda_0)x)V'(x; T) + \lambda_0(V(x \frac{\lambda_1}{\lambda_0}; T) - V(x; T)) = -x$ ,  $V(T; T) = 0$ . If  $\lambda_0 \rightarrow \infty$  and  $\lambda_1 \rightarrow \infty$  such that  $\lambda_1 = \lambda_0 + C\sqrt{\lambda_0}$  for some fixed constant  $C < 0$ , then  $B(T) = B_0(T) + \varepsilon B_1(T) + \dots$ , where  $\varepsilon = 1/\sqrt{\lambda_0}$ ,  $B_0(T) = \frac{1}{\rho} \left[ e^{\frac{1}{\rho T}} \left( -Ei\left(-\frac{1}{\rho T}\right) \right) - \left( 1 - \frac{1}{\rho T} \int_0^\infty \frac{e^{-t} \log(1 + \rho T \cdot t)}{t} dt \right) \right]$ ,  $B_1(T) = \frac{\beta}{\rho^2} \left[ \frac{3}{\rho T} \int_0^\infty \frac{e^{-t} \log(1 + \rho T \cdot t)}{t} dt + e^{\frac{1}{\rho T}} Ei\left(-\frac{1}{\rho T}\right) \cdot \left( \frac{1}{2\rho T} - 1 \right) - \frac{5}{2} \right]$ .

The author is grateful to Professor A.N. Shiryaev for his help.

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## Bayesian Nonparametric Construction of Fleming-Viot Models in Population Genetics

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**Keywords:** Fleming-Viot process; Measure-valued diffusion; Generalised Pölya-urn schemes; Stationary distribution; Two-parameter Poisson-Dirichlet process

**AMS:** Primary: 60G57; 60J35. Secondary: 60J60; 92D25.

### Abstract

The Fleming-Viot process is a probability-measure-valued diffusion which arises as the large population limit of a wide class of population genetics models. In a few of its formulations its stationary distribution is known to be the Dirichlet process or a mixture of Dirichlet processes, but its connections with Bayesian statistics are still to be explored.

This work provides several explicit constructions of Fleming-Viot processes in the Bayesian nonparametric framework, and yields a previously unknown stationary distribution. In particular, by means of known and newly defined generalised Pölya-urn schemes, several types of pure jump particle processes are introduced, describing the evolution in time of an exchangeable population. In each case, the process of empirical measures of the individuals converges in distribution in the Skorohod space to a specific Fleming-Viot diffusion, and the stationary distribution is shown to be the de Finetti measure of the infinite sequence of individuals. In presence of viability selection the stationary distribution turns out to be the two-parameter Poisson-Dirichlet process.

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## Bayesian Transgaussian Kriging

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**Keywords:** Geostatistics; Kriging; Bayesian Statistics

### Abstract

In geostatistics a widely used method for prediction is kriging. It is well known and used for many years. But some limits are inherent on the traditional ways of kriging, i.e. simple, ordinary and universal kriging.

Kriging is based on the assumptions that the covariance function is exactly known and the underlying random field is a gaussian field. In practice, neither the trend or the variogram are exactly known, but on the other hand there may be expert knowledge for the trend or the variogram that should be used for prediction in geostatistics. So a mixture of kriging and Bayesian statistics can be useful.

In theory, Bayesian transgaussian kriging can handle random fields with non-gaussian behaviour and with various trends and different link functions, but the problem is the computational effort. Many applications need near real-time evaluations of the random process, so the CPU-time is limited.

In the article and in the presentation a short overview on the wide field of kriging and Bayesian statistics is given, showing some of the current research the author is doing in moment.

## Session 5

# Stochastic Models in Finance II



## Asymptotic analysis for a simple explicit estimator in Barndorff-Nielsen and Shephard stochastic volatility models

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**Keywords:** Martingale estimating functions, stochastic volatility models with jumps, consistency and asymptotic normality

### Abstract

We provide a simple explicit estimator for discretely observed Barndorff-Nielsen and Shephard models, prove rigorously consistency and asymptotic normality based on the single assumption that all moments of the stationary distribution of the variance process are finite, and give explicit expressions for the asymptotic covariance matrix.

We develop in detail the martingale estimating function approach for a bivariate model, that is not a diffusion, but admits jumps. We do not use ergodicity arguments.

We assume that both, logarithmic returns and instantaneous variance are observed on a discrete grid of fixed width, and the observation horizon tends to infinity. This analysis is a starting point and benchmark for further developments concerning optimal martingale estimating functions, and for theoretical and empirical investigations, that replace the (actually unobserved) variance process with a substitute, such as number or volume of trades or implied variance from option data. This is joint work with Friedrich Hubalek, Technische Universität Vienna.

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## Location as risk factor. Spatial analysis of an insurance data-set

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**Keywords:** Spatial statistics, Claim frequency, Markov random field, Potts model, Monte Carlo EM.

### Abstract

Our aim was to examine the territorial dependence of risk for household insurances. We assumed that the counts of claims of each contractor are Poisson distributed with the appropriate parameter. Besides the classical risk factors such as type of wall, building, etc., we considered the location associated to each contract. In our first approach we used Markov random field model to describe the spatial effect. Basically there are two ways of fitting the model; we fit a GLM to the counts of claims with the classical risk factors and regarding their effects as fixed we fit the spatial model. Alternatively we can estimate the effects of all covariates jointly. Although this latter may seem to be a more accurate approach, its high complexity and computational demands makes it not feasible in our case. To overcome the disadvantages of the separate estimation of the classical and the spatial risk factors we proceeded as follows: first we used a GLM for the non-spatial covariates, and then fitted the spatial model by MCMC. Then we refitted the GLM and the spatial model again, and we iterated this procedure several times. We achieved a much better fit by performing these two steps eight times. Another approach was that instead of using hierarchical Bayes model we estimated the value of the spatial parameters by a maximum likelihood estimate. In this model we assumed Potts model for the relative risk of the regions. As result we achieved analytically insolvable equations, so we used Expectation Maximization (EM) algorithm to get the estimates. In case of high dimension we also used Monte Carlo EM algorithm.

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## Bounded arbitrage for multi-period model of financial market with discrete time

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**Keywords:** Arbitrage.  $\varepsilon$ -arbitrage. Financial market. Multi-period market. Equivalent measure. Fundamental theorem of asset pricing.

### Abstract

In a general discrete time market model with portfolio constraints, an  $\varepsilon$ -arbitrage opportunity is an opportunity to make a riskless profit of amount at least  $\varepsilon$  while trading a limited amount of assets. In a one-period framework with contingent initial data and in a multi-period framework we prove no- $\varepsilon$ -arbitrage criteria similar to the classical fundamental theorem of asset pricing. Classical result is that no-arbitrage is equivalent to the existence of an equivalent martingale measure, i.e. discounted price process is a martingale w.r.t. this measure. We prove that there is no  $\varepsilon$ -arbitrage iff there exists such an equivalent measure with a bounded density that the discounted price process is in certain sense close to a martingale w.r.t. this measure.

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## A discrete model for correlated default times and its application to CDO

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**Keywords:** Default times. Default correlation. Dependence. Portfolio Loss Distribution. CDO.

**AMS:** 62P05

### Abstract

A simple discrete time model for the dependence between the default times of assets (firms) in a given portfolio is discussed. The model is based on multivariate geometric distributions. A possible way for calibration of the model is discussed too. In the framework of this model portfolio loss distribution is calculated and a Monte Carlo based estimation of the expected losses in a given CDO is done.

**Acknowledgements:** This work is partially supported by NFSI grant No.VU-MI-105/2005.

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Session 6

# Stochastic Processes II



## Gaussian Process Simulation with application of the Theory of Square-Gaussian Processes

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**Keywords:** Square-Gaussian processes, simulation, model construction, Gaussian processes, reliability, accuracy

**AMS:** 60G60, 68U20, 65C20, 60G15

### Abstract

In many applied areas which use theory of stochastic processes the problem arises to estimate probability that a random vector process  $\vec{X}^T(t) = (X_1(t), X_2(t), \dots, X_d(t))$  leaves some region on some interval of time. For example, some systems break off on the interval  $[0, T]$  if

$$\sup_{0 \leq t \leq T} \sum_{k=1}^d c_k^2(t) X_k^2(t) > \varepsilon$$

or, in more general case, if

$$\sup_{t \in T} \vec{X}^T(t) A(t) \vec{X}(t) > \varepsilon,$$

where  $\varepsilon$  is a sufficiently large number,  $(T, \rho)$  is a metric space,  $\vec{X} = (\vec{X}(t), t \in T)$  is a process that generates the system,  $A(t)$  is a matrix (in most cases positive semidefinite). The process  $\vec{X}(t)$  may be considered as Gaussian due to the central limit theorem. Thus the problem arises to estimate the probability

$$P \left\{ \sup_{t \in T} \vec{X}^T(t) A(t) \vec{X}(t) > \varepsilon \right\},$$

or the probability

$$P \left\{ \sup_{t \in T} |\vec{X}^T(t) A(t) \vec{X}(t) - \mathbf{E} \vec{X}^T(t) A(t) \vec{X}(t)| > \varepsilon \right\},$$

where  $\vec{X}(t)$  is a Gaussian vector process and  $A(t)$  is a symmetric matrix. The process  $\vec{X}(t)$  is considered as centered one.

The estimates of such large deviation probability for square-Gaussian stochastic processes are obtained. This result is used for model construction of Gaussian stochastic process as input process of same system, taking into account output process, with given reliability and accuracy in Banach space.

## Repairable systems with general repair

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**Keywords:** Repairable system. Virtual age. Hazard Rate. Maintenance policy.

**AMS:** 60K05, 90B25

### Abstract

The contribution deals with general repair models for repairable system. Various ways of modeling the impact of repairs on a system condition are examined. The most common is to assume the repairs impact the failure intensity following a virtual age process (of the general form) proposed by Kijima. Another option considers repairs performing the time-dependent scale transformation of the governing distribution function. Ultimately, repairs directly reducing system deterioration represented by cumulative hazard rate are taking into account. In the second part of the article preventive maintenance actions of repairable system are discussed.

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Session 7

# Resampling



## Resampling methods for the change analysis of dependent data

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**Keywords:** mean change, block resampling, dependent observations

**AMS:** 62G10,62G09

### Abstract

The fundamental question in change-point analysis is whether an observed stochastic process follows one model or whether the underlying model changes at least once during the observational period. This talk focuses on detecting an abrupt mean change in a sequence of dependent data.

In change-point analysis critical values for testing procedures are usually obtained by distributional asymptotics. These critical values, however, do not sufficiently reflect dependency. Moreover it is a well-known fact that convergence rates especially for extreme-value statistics are very slow. Using resampling methods we obtain better approximations, which take possible dependency structures more efficiently into account.

We focus on a block permutation method which splits the observations  $X(1), \dots, X(n)$  into  $L$  blocks of length  $K$ . Instead of resampling  $X(\cdot)$ , we resample the blocks  $X(Kl + 1), \dots, X(K(l + 1)), l = 0, \dots, L - 1$ , but keep the order within the blocks. Then we prove that the original statistics and their resampling counterparts follow the same distributional asymptotics, thus we can use the quantiles of the bootstrap distribution as critical values.

Some simulation studies illustrate that the permutation tests usually behave better than the original tests if performance is measured by the  $\alpha$ - and  $\beta$ -errors, respectively.

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Session 8

# Density Estimation



## Bootstrap Bandwidth Selection Using an $h$ -Dependent Pilot Bandwidth

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**Keywords:** bootstrap bandwidth choice,  $h$ -dependent pilot bandwidth, kernel density estimation, plug-in method

**AMS:** 62G07, 62G20

### Abstract

The problem of choosing the bandwidth  $h$  for kernel density estimation is considered (see [2]). All the plug-in and bootstrap bandwidth selection methods require the use of a pilot bandwidth  $g$  and the usual way to make an  $h$ -dependent choice of  $g$  is by obtaining their asymptotic expressions separately and solving the two equations for the sample size  $n$ . In contrast, we obtain the asymptotically optimal value of  $g$  for every fixed  $h$ , thus making our selection "less asymptotic". Exact error expressions show that some usually assumed hypotheses have to be discarded in the asymptotic study in this case. Based on this idea, two versions of a new bootstrap-type bandwidth selector are proposed, and their properties are analyzed through theoretical results and a simulation study. The new selector can achieve the best possible rate with the minimal asymptotic variance, and the simulations show that in practice one version of this selector provides a good compromise between cross-validation and the popular Sheather-Jones method (see [1]): it is not so variable as cross-validation and its centerpoint is significantly less biased than that of the Sheather-Jones bandwidth selector in most cases.

**Acknowledgements:** This research has been supported by the Spanish Ministerio de Ciencia y Tecnología project MTM2006-06172.

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## Uniform consistency of kernel-type estimators and conditional $U$ -statistics with general bandwidths

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**Keywords:** empirical processes, kernel estimation, Nadaraya–Watson, nonparametric estimation, regression, uniform in bandwidth consistency,  $U$ -statistics.

**AMS:** 60F15, 62G08

### Abstract

Let  $(X, Y), (X_1, Y_1), \dots, (X_n, Y_n)$  be i.i.d. random variables and denote by  $f_X$  the (unknown) marginal density function of  $X$ . For practical use, it is often important to be able to estimate this density, as well as the regression function  $m(t) = \mathbb{E}[Y|X = t]$ . A popular class of estimators consists of kernel-type estimators, which have been intensively studied for a long time. Some well-known examples are the Nadaraya–Watson estimator and the local polynomial regression function estimators.

We start by giving some recent uniform consistency results for kernel-type estimators of  $f_X$  and  $m(t)$ , with a concise overview of the used methodology. The most recent results are shown to be valid uniformly in bandwidth as well, implying that the consistency is preserved when the bandwidth is chosen depending upon the dataset and/or the location. This kind of result is both from theoretical as practical point of view a considerable improvement.

Next, we look at conditional  $U$ -statistics, which were introduced by Stute to estimate the general regression function  $m_\varphi(\mathbf{t}) = \mathbb{E}[\varphi(Y_1, \dots, Y_m)|(X_1, \dots, X_m) = \mathbf{t}]$ , and for which pointwise consistency was proved in 1991. We are now interested in the uniform convergence of these estimators, so a strong uniform consistency result is established, where a uniformity in the bandwidth is considered as well. In addition, the corresponding convergence rates can be derived from this result.

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Session 9

# Statistical Applications I



## Modeling the Mechanism of Postantibiotic Effect

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**Keywords:** Antibiotic resistance. Kolmogorov equations. Penicillin Binding Proteins. Postantibiotic effect

**AMS:** 60J80

### Abstract

Models for the emergence and spread of resistance have many challenges. In this paper we look at the specific problem of modeling the postantibiotic effect (PAE), the delayed regrowth of the bacteria after complete removal of an antibiotic.

A stochastic model for describing one of the possible underlying biological mechanisms of PAE is formulated. The model is based on the theory of penicillin binding proteins (PBPs), where the PAE is the time required by the bacteria to synthesize new PBPs before growth. Newly synthesized PBPs are unsaturated and becomes saturated under antibiotic pressure and eventually removed due to death.

The model assumes that unsaturated PBPs are attached (synthesized) to a bacterium according to a Poisson process and that these in turn are saturated with an intensity proportional to the antibiotic concentration of the treatment. The calculations and results are divided into three simplifying steps toward a more realistic approach. At first, we assume constant antibiotic concentration and no initial PBPs. Secondly, we assume constant antibiotic concentration, but with an initial set of unsaturated PBPs (no saturated PBPs). Thirdly, we assume exponentially declining antibiotic concentration and the same initial set of unsaturated PBPs.

The stochastic models are solved using a set of Kolmogorov equations and exact solutions with interesting properties can be derived for all three steps. The results are useful for giving a better understanding of the time properties of PAE.

## Mixture-Regression Cluster Model applied to Longitudinal Microarray Experiments

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**Keywords:** Microarray; Longitudinal; Mixtures; Regression; Random effects.

**AMS:** 60J80

### Abstract

The aim of this work is to explore various statistical techniques to identify genes which contribute to some change in phenotype level. For example, the response of fish kept under stressful conditions for various lengths of time. We aim to assess the level of *differential* expression of each gene in the tissue samples and also attempt to model the expression patterns of genes over time, not only to classify genes by similarities in expression patterns, but also to model these patterns as specified functions.

The proposed Mixture-Regression Cluster Model is developed to model *and* cluster the genes into groups according to their expressions measured over time. This model is similar to that of the multivariate normal mixture model in that clusters are identified by the EM algorithm but is adapted to incorporate the flexibility of regression curves to fit the trends. In this way, additional features such as covariates, random effects and correlation structures can be incorporated into the model while potentially offering a considerable saving on the number of parameters required to model the trends.

**Acknowledgements:** Special Thanks to colleagues working at the National Diagnostics Centre, and in the Mathematics Department of National University of Ireland, Galway, where the work was carried out.

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## Graphical Modeling for Discrete Random Variables with Application to Tissue Microarray (TMA) Experiments

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**Keywords:** Graphical models, Hierarchical models, Interactions, Lasso, Log-linear models, Variable selection

**AMS:** 68R01, 62P10

### Abstract

Tissue microarrays (TMA) are composed of hundreds of tissue sections from different patients arrayed on a single glass slide. With the use of immunohistochemical staining, they provide a high-throughput method of analyzing potential biomarkers on large patient samples. The assessment of the expression level of a biomarker is usually performed by the pathologist on a categorical scale.

The analysis of the interaction of these biomarkers and in particular the estimation of the graph structure associated with the underlying discrete random variables, are of biological importance. Questions such as how the interaction pattern changes with progressing tumor grade or with survival time are of direct biological interest. However, the estimation of the interaction structure requires sophisticated techniques. Our approach is to fit an  $\ell_1$ -regularized log-linear model assuming a multinomial sampling scheme in order to obtain the graphical model. The regularization becomes necessary as after cross-tabulation of the samples in contingency tables, many cell entries remain zero, leading to so-called sparse contingency tables, where standard procedures fail to work.

We compare our approach with other methods for graphical modeling. Moreover, biological validation of the estimated interaction structure is done by mapping to known biochemical pathways.

## Analysing Regional Firm Startup Activity Using Geographically Weighted Regression: The case of Austria

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**Keywords:** spatial models, geographically weighted regression, founding activity

**AMS:** 91B72

### Abstract

One subject in entrepreneurship research is to find factors which explain regional variations in firm startup activity. Reviewing the entrepreneurship literature points out that in most studies classical linear models are used to identify influencing factors on regional founding rates and that the spatial component of data is mostly disregarded. Estimating a global model assumes that the influence of the independent variables are constant over the whole study region. I will test in this paper if the global linear model is appropriate to estimate the founding activity. Therefor global and local indicators for spatial autocorrelation (Moran I, [1]) are calculated first, to show the spatial dependence in regional firm startup rates and additional regional data on the level of 93 counties of Austria. Further a geographically weighted regression (GWR) model is estimated to test whether a constant relationship between the founding rate and dependent variables exists over the whole study region [2]. The test of spatial dependence results in significant positive autocorrelation for all regional variables. The local parameter values of the GWR model indicate significant regional variations which indicates that a global linear model is not appropriate to explain the startup activity in Austria.

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Session 10

# Markov Processes I



## On asymptotic behavior of a finite Markov chain

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**Keywords:** difference equation; finite Markov chain; infinite product; simulated annealing; uniform weak ergodicity; uniform strong ergodicity; convergence.

**AMS:** 60J10

### Abstract

We consider a finite nonhomogeneous Markov chain with transition matrices  $(P_n)_{n \geq 1}$  such that  $\lim_{n \rightarrow \infty} P_n = P$ , where  $P$  has  $p \geq 1$  irreducible and aperiodic closed classes and, perhaps, transient states. We give sufficient conditions for uniform weak and uniform strong ergodicity and convergence, respectively, of the nonhomogeneous chain in terms of similar properties of a nonhomogeneous Markov chain of smaller size. We show the dependence of long-run behavior of  $(P_n)_{n \geq 1}$  on spectral properties of the matrix  $P$  and on the nature of the perturbations  $V_n = P_n - P$ ,  $\forall n \geq 1$ . Our conditions are compared to some other results existing in literature. Markov chains of this type occur in simulated annealing, a stochastic algorithm for global optimization.

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## Criteria for transience of branching Markov chains

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**Keywords:** branching Markov chains, transience, recurrence, superharmonic functions

**AMS:** 60J10, 60J80

### Abstract

A branching Markov chain (BMC) is a system of particles in discrete time. The BMC starts with one particle in an arbitrary starting position  $x$ . At each time particles split up in offspring particles independently according to some probability distributions that may depend on the locations of the particles. The new particles then move independently according to a Markov Chain.

An irreducible Markov Chain is either recurrent or transient: either all or none states are visited infinitely often. It turns out that this dichotomy breaks down for BMC and that one can classify BMCs in three different types. Let  $\alpha(x)$  be the probability that, starting the BMC in  $x$ , the state  $x$  is hit infinitely often by some particles. There are three possible regimes: *transient* ( $\alpha(x) = 0 \forall x$ ), *weakly recurrent* ( $0 < \alpha(x) < 1 \forall x$ ) and *strongly recurrent* ( $\alpha(x) = 1 \forall x$ ). We give equivalent criteria for transience of BMC and discuss some interesting consequences.

## A Regeneration Proof of the Central Limit Theorem for Uniformly Ergodic Markov Chains

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**Keywords:** Markov Chain, CLT, Uniform Ergodicity, Regeneration.

**AMS:** 60J05

### Abstract

Central limit theorems for functionals of general state space Markov chains are of crucial importance in sensible implementation of Markov chain Monte Carlo algorithms as well as of vital theoretical interest. Different approaches to proving this type of results under diverse assumptions led to a large variety of CLT versions. However due to the recent development of the regeneration theory of Markov chains, many classical CLTs can be reproved using this intuitive probabilistic approach, avoiding technicalities of original proofs. In this paper we provide a regeneration proof of a CLT for functionals of uniformly ergodic Markov chains, thus solve the open problem posed in [3]. Moreover we discuss the difference between one-step and multiple-step small set condition.

The presentation is joint work with Witold Bednorz from Warsaw University.

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## On the statistics of branching processes with a random number of ancestors

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**Keywords:** offspring mean, individual distribution, robust estimation

**AMS:** 60J80

### Abstract

In the present paper we consider the robust estimation in the Bienayme - Galton - Watson processes with a random number of ancestors (BGWR processes) in the sense of the weighted and trimmed likelihood. The classical estimators of the individual mean and distribution are compared with their robust modifications via simulations and computational results.

**Acknowledgements:** This research was partially supported by the National Science Fund of Bulgaria, Grant No VU-MI-105/2005.

Session 11

# Extreme Value Theory



## Asymptotic and pre-asymptotic tail behaviour of a power max-autoregressive model

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**Keywords:** Markov Chains. Extreme value theory. Stationary sequence. Dependence conditions. Tail index. Extremal index. Tail dependence index. Asymptotic independence.

**AMS:** 60G70

### Abstract

Nowadays we can not ignore the demand of various areas like hydrology, geophysics or finances, on modeling extreme data or exceedance data above certain high levels. The assumption of independent observations gave place to a more realistic statement: the temporal dependence. Phenomenon therein are often associated with markovian sequences for which we claim simple ways of modeling in what concerns the extremes topic, like for instance, *max-autoregressive* processes (Alpuim [1] and Canto e Castro [3]). Here we present a max-autoregressive process involving a power transformation -  $ARMAX_p$  - whose parameter  $c \in (0, 1)$  relates directly with the coefficient of tail dependence ( $\eta$ ) of Ledford and Tawn [4]. This index characterizes the penultimate tail dependence of a process and can be related with a threshold-dependent extremal index, which assumes an important role when extending discussions of extreme values from independent and identically distributed (i.i.d.) sequences to stationary ones (Bortot and Tawn, [2]). In this paper, we analyze some asymptotic extremal features, we study the local dependence structure and we state the threshold-dependent extremal index connection with parameter  $\eta$ , for the  $ARMAX_p$  process.

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## Estimators of the extreme value index based on quantile regression

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**Keywords:** extreme value index, quantile regression

**AMS:** 60G70

### Abstract

In the contribution we discuss an estimation of the extreme value index in the case of linear regression model. We propose estimators based on the extreme regression quantiles slightly expanding older results of Chernozhukov [1]. Our aim is to describe a general class of estimators of the extreme value index based on regression quantiles. We show consistency and asymptotic normality of estimators belonging to this class using methodological tools based on Dress (see [2], [3]) and Gutenbrunner [4]. Various conditions and their necessity are discussed. We demonstrate a performance of proposed estimators on a small simulation study.

**Acknowledgements:** This work was supported by Czech Republic Grant GAČR 201/05/H007.

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Session 12

# Markov Processes II



## MCMC Sampling for Diffusion Processes

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**Keywords:** Stochastic Differential Equation, Markov chain Monte Carlo, Particle Filter

### Abstract

It is well known that likelihood inference for arbitrary nonlinear diffusion processes observed at discrete times is problematic since closed form transition densities are rarely available. One widely used solution adopts the treatment of Pedersen [1] and involves the introduction of latent data points between every pair of observations to allow an Euler-Maruyama approximation of the true transition densities to become accurate. In recent literature, Markov chain Monte Carlo (MCMC) methods have been used to sample the posterior distribution of latent data and model parameters; however, naive schemes suffer from a mixing problem, highlighted by Roberts and Stramer [2], that worsens with the degree of augmentation. We will consider some recently developed MCMC schemes that are not adversely affected by the amount of augmentation. The methodology will be illustrated by estimating parameters governing the diffusion approximations of some interesting systems biological models.

**Acknowledgements:** This research is joint work with Darren J. Wilkinson, School of Mathematics & Statistics, Newcastle University, UK

### References

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## Recent advances in the simulation and inference for stochastic differential equation models

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**Keywords:** Exact simulation, Monte Carlo maximum likelihood, filtering, EM algorithm

### Abstract

In this talk I will review recent advances in the exact simulation of diffusion processes, introduced mainly in [1], [2] and [3]. I will present a general methodology which can be used to solve a variety of statistical problems related to diffusion processes and I will give special attention to an EM algorithm for the estimation of parameters and unobserved states.

### References

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## Estimating equations and inference from diffusion driven models

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**Keywords:** general estimating function, goodness of fit testing, stochastic differential equation.

### Abstract

Diffusion-type models provide a natural and flexible framework for modeling phenomena that evolve continuously and randomly with time. However their statistical analysis is complicated. Only rarely the functional form of the likelihood function of a diffusion is explicitly known. The same problem of course occurs with diffusion driven models such as summed diffusions, integrated diffusions, and stochastic volatility models. General estimating equations often present a simple alternative means for fitting these models. I will illustrate the use of estimating equations for diffusion-type models driven by diffusions having linear drift and quadratic squared diffusion coefficient. I shall use the term Pearson diffusions as the invariant distributions of this class of diffusions belong to the Pearson system. Further I will present a new goodness of fit test based on estimating equations for varying sampling frequencies. The basic idea is to compare the parameter estimates based on the full data to those obtained from downsamples. The test can be used to distinguish different kinds of diffusion driven models.

**Acknowledgements:** The research presented here is joint work with Michael Sørensen, Department of Mathematical Sciences, University of Copenhagen and with Bo Markusen and Helle Sørensen, Department of Natural Sciences, LIFE, University of Copenhagen.

## Examples on lag distributed models subject to nonnegative divided differences of orders 2, 3 and 4

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**Keywords:** approximation, divided differences, lag distributed models, least squares, quadratic programming, r-convexity.

**AMS:** 60J80

### Abstract

We consider noisy measurements from a time series that follow a linearly distributed lag model. It is usual to assume that the lag coefficients lie on some curve and then specify the curve by a least squares calculation. However, we define the  $r$ -th order smoothness priors by requiring nonnegative divided differences of order  $r$  for the lag coefficients. Such priors do not imply any parameterization of the lag curve and provide a more accurate representation of the prior knowledge. For the calculation of the solution we propose an algorithm that gives the least squares change to the data subject to nonnegative divided differences of the lag coefficients of order  $r$ , where  $r$  is a prescribed positive integer. The problem is a strictly convex quadratic programming calculation, where each of the constraints functions depends on  $r+1$  adjacent components of the smoothed values of the lag coefficients. We take account of this special structure and use a special active set method that is more efficient than general quadratic programming algorithms. In fact we construct a basis that reduces the equality-constrained minimization calculations that appear during the quadratic programming iterations to unconstrained minimization ones, which depend on much fewer variables. Finally, we present an example that illustrates our approach.

**Acknowledgements:** The work of the author is supported by research grand 70/7/7145 (Heracleitus) of the university of Athens, which is co-financed within Op. Education by the European Social Fund and National Resources. We feel grateful to them all.

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Session 13

# Hypothesis Testing II



## Some multivariate multi-sample tests based on spatial medians

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**Keywords:** Spatial median. Spherical and elliptical symmetry. Asymptotic equivalence.

**AMS:** 62H11

### Abstract

The multivariate multi-sample location problem in case of spherically or elliptically symmetric distributions is considered and we test the hypothesis about equality of the location parameters (see e.g. [1]). We present some analogies of the well-known Lawley-Hotelling test statistic based on spatial medians of the samples.

In case of spherical symmetry, asymptotic distribution of the median-based test statistics is derived under the hypothesis and also under the sequence of Pitman's alternatives. Using the Bahadur-type representation of the spatial median (see [2]) the asymptotic equality of the median-based statistics to an obvious multi-sample extension of the well-known Rayleigh test statistic is proved under the hypothesis.

Some Monte Carlo results are introduced, involving also the case when the underlying distribution is elliptically symmetric.

**Acknowledgements:** The research was supported by the VEGA grant 1/3016/06 of the Scientific Grant Agency of the Slovak Republic.

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## Approximations to Most Powerful Invariant Tests for Multinormality

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**Keywords:** Most powerful invariant test. Laplace approximation.

**AMS:** 62H15

### Abstract

We consider the problem of testing multinormality against alternatives invariant with respect to some subgroup of affine transformations. In [3], a general form of the most powerful invariant (MPI) test has been obtained. Unfortunately, applicability of the MPI test is rather limited, due to complicated, intractable integrals. With the aid of the Laplace method for integrals, we derive large sample approximations for the MPI tests. The cases of bivariate exponential and uniform alternatives are studied in details, whereas higher dimensional extensions are outlined. It is shown in the both bivariate cases, that a further approximation for the Laplace approximation can be given. This leads to the likelihood ratio (LR) test statistic. A final conclusion is that the likelihood ratio test statistic can be seen as a formal expansion of the MPI test statistic, with a known upper bound for the relative error of the approximation. The Monte Carlo simulation study shows, that powers of both, the Laplace approximation, as well as the LR test are very close to the power of the most powerful invariant test even in small sample sizes.

**Acknowledgements:** The research was partly supported by the KBN local grant No 11.420.04

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Session 14

# Random Matrices



## Large dimensional random matrices

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**Keywords:** Sample covariance matrix. eigenvalues and eigenvectors. wireless communication

**AMS:** 15A52;60F15

### Abstract

This paper focuses on the theory of spectral analysis of Large Random Matrix. Concerning eigenvalues and eigenvectors some important results and methods are reviewed and moreover, some latest results regarding eigenvectors are also presented. Finally, the application of random matrix to wireless communication is illustrated.

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Session 15

Time Series



## Principal Component Analysis of Dependent Data

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**Keywords:** Principal Component Analysis. Time Series. SSA. MSSA.

**AMS:** 62H25, 62M15

### Abstract

In many of the statistical methods, and in particular in Principal Component Analysis (PCA), it is assumed that all  $n$  observations  $x_1, x_2, \dots, x_n$  are independent. The aim of this work is to analyze the consequences of the application of PCA to observations that do not verify this hypothesis.

Skinner [5] and Konishi [2], among others, state that the type of dependence in the observations is varied. Being time series the most frequent situation in practice, we present the Singular Spectrum Analysis (SSA), Golyandina [1], and the Multichannel Singular Spectrum Analysis (MSSA), Oliveira [4] or Kim [3] as methods based in PCA.

This work is divided in two parts. In the first we present a description of SSA, which uses the auto-covariance matrix of a time series with only one variable and one application. In the second part we present the multivariate case, the MSSA, together with another application.

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## Comparison of prediction quality of the best linear unbiased predictors in time series linear regression models.

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**Keywords:** Time series. Finite discrete spectrum linear regression models. Best linear unbiased prediction. Mean squared error of a predictor. Löwner partial ordering.  
**AMS:** 62M10, 62M20

### Abstract

Using a matrix-block version of the mean squared error of the best linear unbiased predictor (BLUP) and the Löwner partial order relation for symmetric matrices, we derive and present our results dealing with a comparison of prediction quality of BLUP in different types of time series linear regression models which belong to a general class of the so-called finite discrete spectrum linear regression models (FDLSRM). Some of the results are completely new and some of them are generalization of recent results achieved in special orthogonal cases of FDLSRM.

**Acknowledgements:** This work was supported by the grant 1/3001/06 of the Slovak Scientific Grant Agency (VEGA) and the grant 25/07-08 of VVGS at P. J. Šafárik University in Košice.

Session 16

# Statistical Applications II



## Recent Advances in Palaeoclimate Reconstruction

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**Keywords:** Bayesian statistics; Spatial Models; Paleostatistics  
**AMS:**

### Abstract

This talk will concentrate on some of the developments since Haslett et al. [1]. The premise is that climate change affects pollen distribution; fossil pollen from lake sediment thus gives information on historical climate. By building a training data set which links modern pollen and modern climate, we can calibrate fossil pollen into historical climate. The Bayesian formulation of this problem contains many interesting facets, including spatial modelling in high dimensions, radiocarbon dating of depth chronologies, and some aspects of zero-inflation. In particular, the talk will discuss how smoothness in climate change is modelled, and extensions which aim to reconstruct the entire climate of Western Europe.

### References

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## On spatial extremes: with application to a rainfall problem

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**Keywords:** spatial extremes, max-stable process, areal reduction factor.

**AMS:** 60G70; 62G32; 62M30;

### Abstract

We consider daily rainfall observations at 32 stations in the province of North Holland (The Netherlands) during 30 years. Let  $Q$  be the *total* rainfall in this area on one day. An important question is: what is the amount of rainfall  $Q$  that is exceeded once in 100 years? This is clearly a problem belonging to extreme value theory. Also it is a genuinely spatial problem.

Recently, a theory of extremes of continuous stochastic processes has been developed. Using the ideas of that theory and much computer power (simulations) we have been able to come up with a reasonable answer to the question above.

Session 17

# Theory of Codes



## Probabilistic models in quasigroup error-detecting codes

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**Keywords:** probabilistic models, error-detecting codes, quasigroup.

**AMS:** 94B70

### Abstract

In my future work I plan to investigate error-detecting codes and probabilistic models in them. Error-detecting codes are used to detect errors when messages are transmitted through a noisy communication channel. We propose error-detecting codes using quasigroup operation  $*$  on set  $A = \{0, 1, 2, 3\}$ . In order to detect errors, we extend an input message  $a_1a_2 \dots a_n$  to block  $a_1a_2 \dots a_nd_1d_2 \dots d_n$ , where  $d_i = a_i * a_{(i+1) \bmod n}$ ,  $i = 1, 2, \dots, n$ . We calculate an approximate formula which gives the probability that there will be errors which will not be detected.

**Acknowledgements:** These problem was formulated by professor Verica Bakeva, Institute of Informatics, Faculty of Natural Sciences and Mathematics, Skopje.

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## Random codes

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**Keywords:** random codes, probabilistic models.

**AMS:** 94B60

### Abstract

The non-constructive proof of the noisy-channel coding theorem shows that good block codes exist for any noisy channel, and moreover that nearly all block codes are good. However, writing down an explicit and practical encoder and decoder that are as good as proved by Shannon in his seminal work *A Mathematical Theory of Communication* is still an unsolved problem. Recently, it has been recognized that two classes of codes, namely turbo codes and low-density parity-check (LDPC) codes, perform at rates extremely close to the Shannon limit. Turbo and LDPC codes are based on a similar philosophy: constrained random code, described by some fixed parameters plus randomness, decoded using iterative algorithms or message passing decoders. My plains for future work are investigations to probabilistic models in coding theory, especially in random codes.

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